

Finite-sample quantiles of the Jarque–Bera test

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The finite-sample null distribution of the Jarque–Bera Lagrange multiplier test for normality differs considerably from the asymptotic $\chi^2(2)$. However, asymptotic critical values are commonly used in applied work, even for relatively small sample sizes. Here, very accurate response surface approximations are developed for the 10% and 5% critical values of the test, which enable correct practical implementation.

I. Introduction

The Jarque–Bera (1980, 1987) Lagrange multiplier test is perhaps the most commonly used procedure for testing whether a univariate sample of T datapoints, or estimated regression residuals, are drawn from a normal distribution. It is a joint test of the null hypothesis (of normality) that sample skewness equals 0 and sample kurtosis equals 3, and the null is rejected when the statistic

$$LM = \frac{T}{6} \left\{ (b_1^{1/2})^2 + \frac{(b_2 - 3)^2}{4} \right\} \stackrel{a}{\sim} \chi^2(2)$$

exceeds some critical value, which is usually taken from the asymptotic $\chi^2(2)$ distribution. The standardized third and fourth moments are given by $b_1^{1/2} = (m_3/m_2^{3/2})$ and $b_2 = (m_4/m_2^2)$ respectively, and m_i is the i th central moment of the sample. It has been noted that the small-sample tail quantiles of the LM statistics are quite different from their asymptotic counterparts: e.g. Deb and Sefton (1996, Table 1) and Urzúa (1996, Table 1). The use of asymptotic critical values given even fairly large samples will distort the actual size of the test, and may lead to incorrect decisions in applied work.

Deb and Sefton (1996) compute 14 very accurate empirical 10% and 5% significance points of LM in the interval $T \in [20, 800]$, and show that their use gives an almost correctly-sized test using regressors. However, practical implementation using their critical values requires new simulations for sample sizes that are not tabulated. We address this problem, and develop highly accurate response surface¹ approximations to the 10% and 5% finite-sample critical values of LM, that are generally correct to ± 0.01 , and may be used for $T \geq 5$.

II. Finite-Sample Critical Values

Using Monte Carlo simulation, we generate 1 000 000 realization of LM under the null of normality, for each sample size T in the set

$$T \in \{5, 6, \dots, 25, 30, \dots, 100, 125, \dots, 1000\}, \quad (1)$$

$$\alpha \in \{0.90, 0.95\}$$

and calculate the 10% and 5% critical values as the $\alpha(1\,000\,000)$ th largest values of LM. This procedure gives 72 datapoints for each α , which are rather more accurate than those previously available. We generate standard normal pseudorandom numbers similarly

¹Response surfaces are numerical-analytical approximations, that have been widely applied in econometrics; e.g., Ericsson (1991), Cheung and Lai (1995), MacKinnon (1994), MacKinnon *et al.* (1999), and Ericsson and MacKinnon (2002).

to MacKinnon's (1994, p. 170) long-period algorithm. The sample sizes that we have chosen are representative of those that are commonly used in applied work. Our design focuses on $T \leq 25$, since the actual critical values vary widely across this range, and specification of good response surfaces requires this information. All simulations were performed on a Pentium 4 machine, with a 2GHz

processor and 256 MB of RAM, running GAUSS under Microsoft Windows XP.

We regressed the Monte Carlo estimates of $100\alpha\%$ quantiles on various functions of sample size, constructed so that $T \rightarrow \infty$ gives the (known) asymptotic quantiles. Following much experimentation, and motivated by the quantile approximations developed by MacKinnon *et al.* (1999, p. 569) in the context of Johansen-type tests for cointegration, we chose to fit the following quantile response surface:

Table 1. Estimated quantile response surfaces \hat{q}^α

| | $\alpha = 0.90$ | $\alpha = 0.95$ |
|-------------------|-----------------|-----------------|
| q_∞^α | 4.605049 | 5.9913104 |
| β_1 | -145.1816602 | -67.00449919 |
| β_2 | 7286.233799 | 1719.108744 |
| β_3 | -275 153.9753 | -74.443.10488 |
| β_4 | 6 437 304.253 | 1962 801.944 |
| β_5 | -92 456 006.82 | -30 095 541.45 |
| β_6 | 814 503 598.1 | 275 285 058.6 |
| β_7 | -4 276 230 401 | -1 479 198 621 |
| β_8 | 12 243 649 840 | 4 299 485 882 |
| β_9 | -14 677 406 860 | -5 206 421 393 |
| R^2 | 0.9999 | 0.9999 |
| RSS | 0.006423 | 0.009567 |
| Mean $ \hat{u} $ | 0.00782 | 0.00899 |
| Max $ \hat{u} $ | 0.0230 | 0.0344 |

Notes: The response surfaces (2) were estimated in E-Views. All estimated coefficients were significant at the 1% level. R^2 is the degrees-of-freedom adjusted coefficient of determination. RSS is the residual sum of squares. Mean $|\hat{u}|$ is the mean absolute error of the response surface approximations against the simulated critical values, and Max $|\hat{u}|$ is the maximum absolute error.

$$q^\alpha(T_i) = q_\infty^\alpha + \sum_{k=1}^9 \beta_k T_i^{-k} + u_i \quad (2)$$

The dependent variable $q^\alpha(T_i)$ is the simulated finite-sample $100\alpha\%$ quantile with sample size T_i , which takes values from (1); q_∞^α is the asymptotic $100\alpha\%$ quantile from the $\chi^2(2)$ distribution, which was computed in GAUSS as $q_\infty^\alpha = \arg \min\{\text{cdfschic}(q,2) - \alpha + 1\}^2$; u_i is an error term.

We denote the estimated response surface by \hat{q}^α , and estimated coefficients are reported in Table 1. Selection criteria included small residual variance, parsimony, and satisfactory diagnostic performance. The dependencies of the 10% and 5% critical values on sample size are presented in Figs 1–3, which plot response surfaces \hat{q}^α against T . The response surface fits are very good, and generally agree with the simulated quantiles to roughly ± 0.01 , across the entire parameter space (1). For instance, estimated critical values are 2.75 (10%) and 4.41 (5%) for a sample size of 30, and 3.48 (10%) and 5.28 (5%) for a sample size of 75. Simulated critical values are 2.74 (10%)

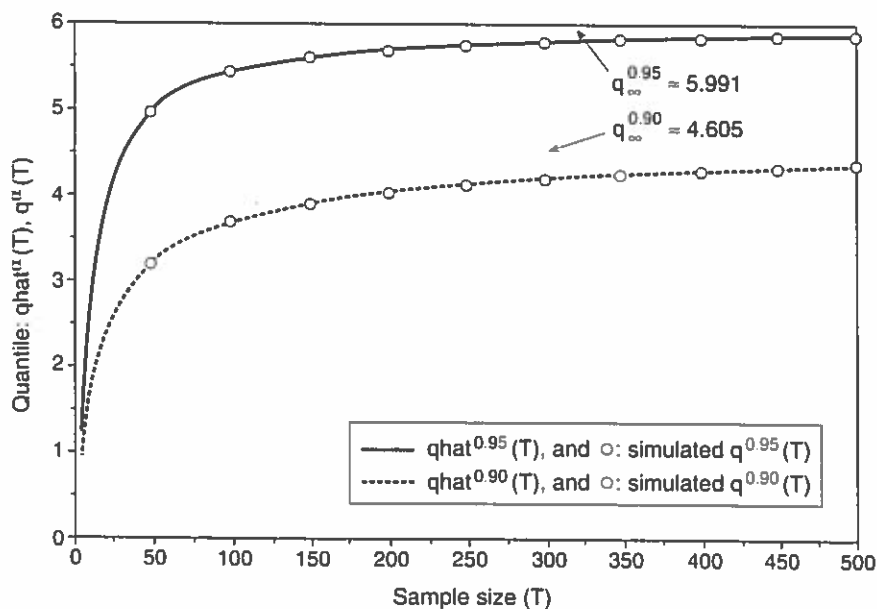


Fig. 1. Quantile response surfaces $\hat{q}^{0.95}$ and $\hat{q}^{0.90}$, for $T \in [5,500]$

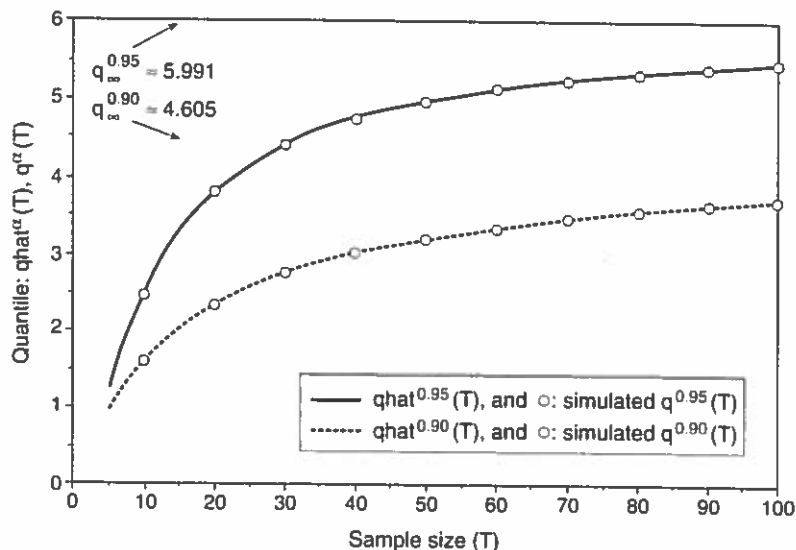


Fig. 2. Quantile response surfaces $\hat{q}^{0.95}$ and $\hat{q}^{0.90}$, for $T \in [5, 100]$

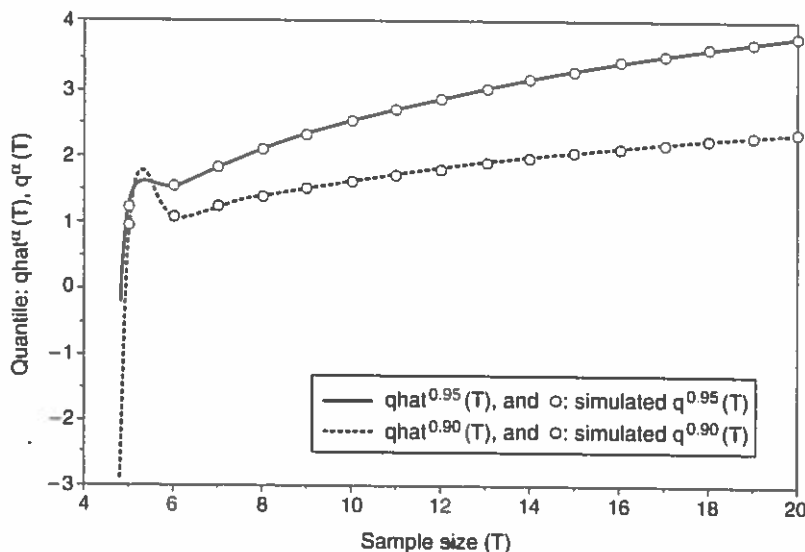


Fig. 3. Quantile response surfaces $\hat{q}^{0.95}$ and $\hat{q}^{0.90}$, for $T \in [5, 20]$

and 4.41 (5%) for a sample size of 30, and 3.49 (10%) and 5.27 (5%) for a sample size of 75. We considered more parsimonious approximations, with fewer inverse powers of T , although these failed to yield an improved fit over (2). Clearly, $\hat{q}^{0.90}$ and $\hat{q}^{0.95}$ break-down for $T \leq 4$, although this is unlikely to be a problem in applied work.

III. Concluding Comments

We have developed very accurate response surface approximations to the 10% and 5% critical values

of the Jarque-Bera test for normality, that may be used to give an almost correctly-sized test in empirical work. While other tests for normality are also available (see Lobato and Velasco, 2004, and references therein), we focus on Jarque-Bera due to its great popularity in applied econometric studies.

Various authors have proposed Monte Carlo (e.g., bootstrap) methods for correcting the finite-sample size distortions of residual-based tests, e.g. Dufour *et al.* (1998) and Kilian and Demiroglu (2000). However, the approach used in this paper is very simple to implement, inexpensive when compared to more computational-based techniques for inference.

and can easily be generalized to other tail quantiles if necessary.

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